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Effective Implementation of Bayesian Adaptive Randomization in Early Phase Clinical Development

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Joint work with

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- Oncology Proof Of Concept Trials: Some Considerations
- Bayesian Adaptive Randomization methodology
- Case Study
- Summary of Simulation Results
- Discussion and Conclusions





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OBJECTIVES:

- Activity: determine whether the treatment is sufficiently promising to proceed in further development
 Safety: better characterize the safety profile of the compound
 Doses: determine the best dose (efficacy / safety)
- Biomarkers: for stratification or prediction of response
- **Strategy**: Add-on strategy or replacement strategy

Challenging design and studies given their limited size and duration !

SINGLE ARM STUDIES

- Endpoint: Response Rate or rates of PFS/OS at predefined timepoint
- Early stopping rules for futility (Simon two-stage design)
- Designed for cytotoxic compounds, not fitting with compounds with different Mode Of Action
- Designs characteristics not consistent with phase III program
 - Not comparative with efficacy hypothesis testing based on historical control
 - Endpoints not used in phase III programs
 - Selection bias
- Difficult assessment of add-on therapies

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SCREENING DESIGNS

- Design characteristics similar to phase III studies
 - Time To Event Endpoints used (PFS more frequently than OS)
 - Comparative
 Treatment effect (HR)
 Hypothesis testing procedure (Log-rank)
 - Randomized
 Selection bias better controlled
- Sample Size smaller than phase III trials but wider than single arm studies (150 / 300 subjects)
 - ➡ Inflation of type I and II error rates ➡ alpha 10% 30%; power ~ 80%
 - Not optimal decision making process
 - Limited to address dose-response or biomarker questions

MAIN CHALLENGES

- Learning phase of development
 still limited knowledge on compound characteristics during study planning
- Classical study designs
 - Fixed treatment allocation
 - No changes allowed during the trial
 - Design independent of data generated during the study
- In studies of limited size, many subjects exposed to control may not be informative (e.g. for safety or for predictive biomarkers)

Overview



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- Bayesian Adaptive Randomization
 - Concept & Rationale
 - Workflow
 - Statistical model
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- CONCEPT
 - Trial design: randomized & comparative
 - Adapt the randomization ratio during the study favoring treatment arm(s) showing best performance
 - Intermediate data of activity available during the study will be used to perform the adaptation
 - Implement efficient stopping rule for futility as soon as the drug shows no activity

Bayesian Adaptive Randomization

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- Fewer subjects assigned to less effective treatment arms
- Keep flexible design during a learning / exploratory phase of development
- Use prior information on the compound and specific indication setting (Bayesian)
- More information on experimental treatment arm (if active)
 - increased precision in the point estimates of activity within arm
 - more safety information
 - improve dose selection
- Improve decision making process

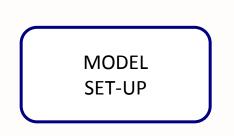
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Bayesian Adaptive Randomization: workflow





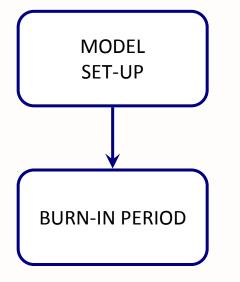
Step 0: Preliminary activity before start of the study

- Feasibility of the design
- Definition of prior information to be included in the model
- Fine tuning of model parameters
- Evaluation of operating characteristics versus standard designs

TOOL: SIMULATIONS

Bayesian Adaptive Randomization: workflow





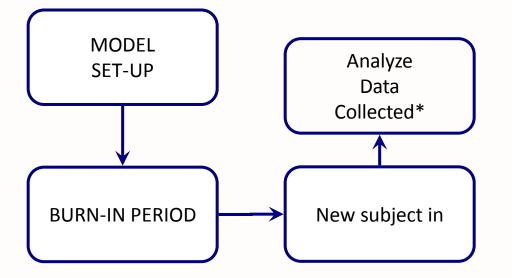
STUDY START:

Step 1: BURN-IN PERIOD

- First group of subjects is assigned to treatment arms according to standard procedures (block randomization with equal allocation ratio)
- Allows model to incorporate enough information to adapt the randomization in a robust way

Bayesian Adaptive Randomization: workflow



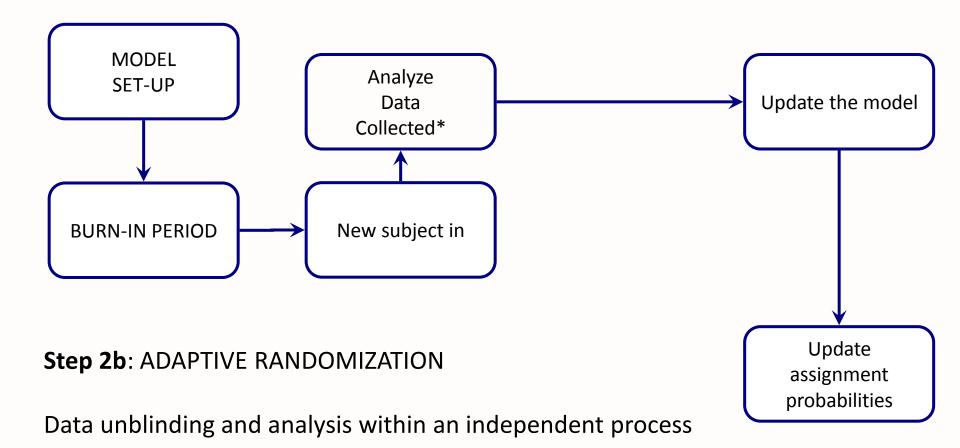


Step 2a: ADAPTIVE RANDOMIZATION

At the completion of the burn-in period before new subject is randomized

Data are transferred from the clinical database to IVRS supplier

Bayesian Adaptive Randomization: workflow

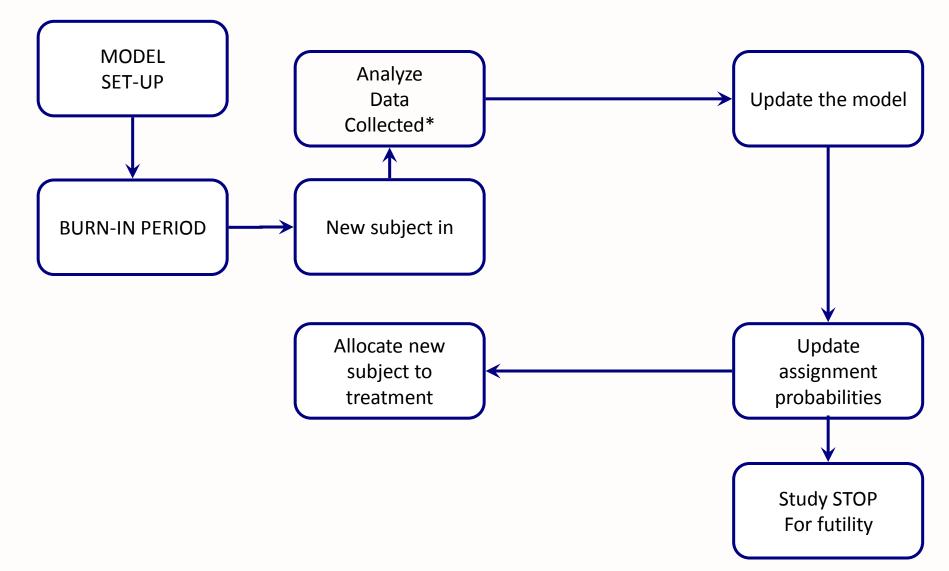


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Trial Team and sponsor blinding should be adequately insured

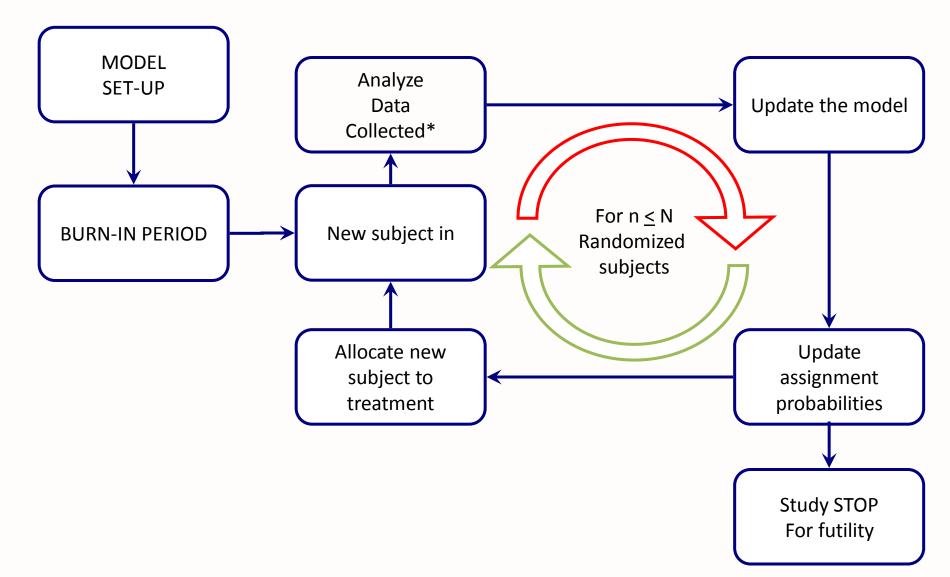
Bayesian Adaptive Randomization: workflow





Bayesian Adaptive Randomization: workflow





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- Assignment probabilities are derived by combining prior information with observed data
 - Guarantees that observed likelihood does not exclusively drive the adaptive randomization.
- Prior information
 - Summarizes previous knowledge on the control arm (literature data) and on the experimental treatment arm (previous trials / preclinical / expectations)
 - Priors should not favor the experimental arm
 bias the randomization process.

Model specification: the statistical engine Cutel

Evolution of the randomized "play the winner" design

Model links the chance of assigning a subject to one treatment arm [γ] to the probability that that treatment has the best performance over the other(s) [π]

 $\gamma_i(i)$ = Probability [subject *i* is randomized to treatment *j*]

$$\pi_i = \text{Prob}(\eta_i > \max(\eta_k) \mid \text{data, prior}) \quad \text{for } k \neq j$$

Posterior Prob [primary endpoint in treatment *j* > all other arms]

 $\gamma_{j}(i) = \pi_{j}(i)^{\lambda} / \Sigma_{j}\pi_{j}(i)^{\lambda}$

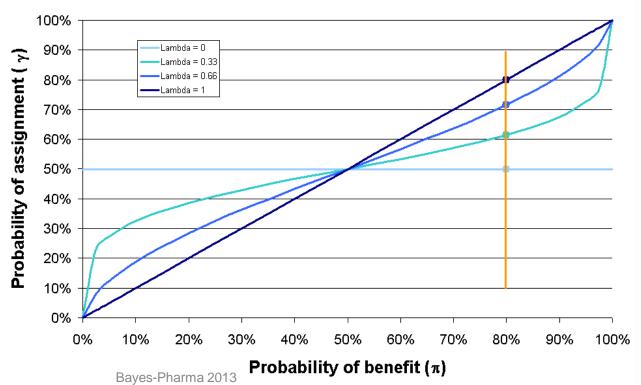
Model specification: the statistical engine Cytel

 $\gamma_j(i) = \pi_j(i)^{\lambda} / \Sigma \pi(i)^{\lambda}$

 λ = tuning parameter controlling the degrees of freedom of the process

- $\lambda = 0 \Rightarrow$ balanced randomization
- $\lambda = 1 \Rightarrow \gamma_j(i) = \pi_j(i)$

The value of lambda based on simulation results before study start





π_{soc} = Probability (Standard Of Care > Experimental Treatment Arm(s))

Direct measure of drug activity to be used for decision making

• During the study High $\pi_{SOC} \ge c_1$

Stop for futility for weak drug activity

• Final analysis Low $\pi_{Soc} \le c_2$

Claim drug activity within a hypothesis testing framework

 Simulation results will pre-define proper values for c₁ and c₂ leading to adequate control of type I and II error





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Trial Insights



- Cytostatic compound (monoclonal antibody) with not established doseresponse curve (monotonic or bell-shaped)
- Phase II randomized
 - Standard Of Care (SoC)
 - SoC + "LOW" dose
 - SoC + "HIGH" dose
- Primary endpoint: Progression Free Survival
- Study Objective
 - **Primary**: Evaluate Drug activity
 - Secondary: Choose the best dose



Standard solution not completely satisfactory as

- Two parallel looks to data lead to multiplicity issues inflating alpha and increasing the power
 - the overall false positive rate (alpha) equal to 23%
 - power > 90% in case both arms are equally active
- Not feasible to have clearer and more robust decision rule for selection of the best dose
- Performance of Bayesian Adaptive Randomization evaluated through simulations

Trial Insights: Simulation plan



- Simulations were run to
 - Evaluate model operating characteristics versus standard design
 - Define the model parameters (burn-in period, tuning parameter, priors, futility stopping rule, rejection region definition)
- Scenarios of activity:
 - Negative
 - No drug activity
 - One arm active (mild / strong activity)
 - Dose response (mild / strong activity)
 - Both arms equally active (mild / strong activity)

Simulation scenaria

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Scenario 1					
Negative	Hypoth	esis	$\lambda_{P} > \lambda_{L} = \lambda_{H}$		
	Median	(months)	2.76	2.30	2.30
	HR		120%	120%	
	λ	3.98	3.32	3.32	
Scenario 2					
Null	Hypoth	esis	$\lambda_{\rm P} = \lambda_{\rm L} = \lambda_{\rm H}$		
hypothesis	Median	(months)	2.76	2.76	2.76 Used for
	HR		100%	100%	estimation of typ
	λ	3.98	3.98	3.98	l error
Scenario 3.	1				
Only one	Hypoth	esis	$\lambda_{P} = \lambda_{L} < \lambda_{H}$		
arm	Median(months)		2.76	2.76	3.68
moderately	/ HR		100%	75%	
active	λ	3.98	3.98	5.31	
					Used for
Scenario 3.2	2				estimation of
Only one	Hypoth	esis	$\lambda_{P} = \lambda_{L} < \lambda_{H}$		
arm highly			2.76	2.76	4.14 power
active	HR	,	100%	66.7%	
	λ	3.98	3.98	5.97	





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Summary of results: Model parameter



- **Burn-in period**: 81 subjects randomized (37% of sample size)
- Dynamic tuning parameter = (2/3) * (number or subjects randomized / max sample size)
- Futility rule: study stopped anytime for futility when
 - $\pi_{SoC} = P[SoC > experimental arms] > 0.6$ and
 - \geq 150 subjects randomized (2/3 of sample size)
- Null Hypothesis of "No drug effect" rejected \Rightarrow if $\pi_{Soc} < 0.095$



Choice of the Priors

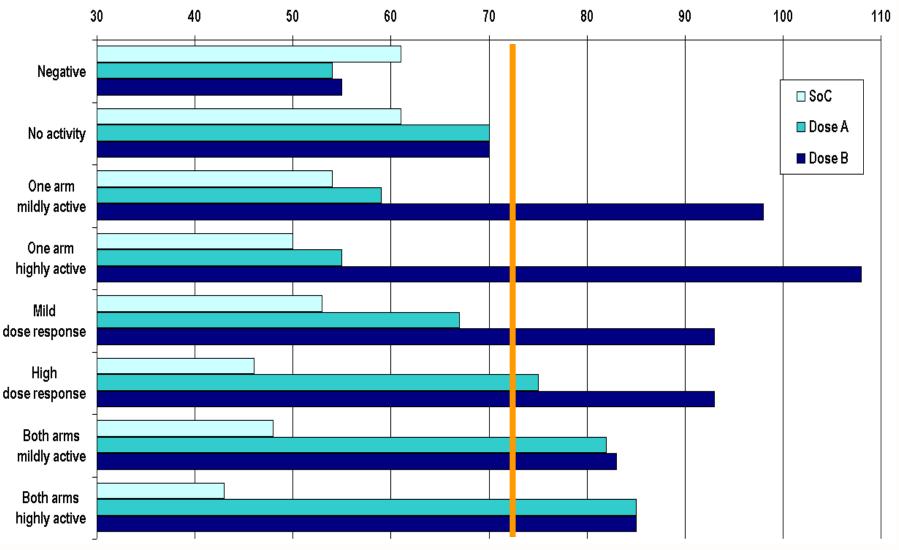
- Standard Of Care
 - Point estimate and 95% confidence interval of most recent and relevant pivotal studies in the same setting.
- Experimental treatment
 - Same expected point estimate
 no drug activity assumed in order not to bias the randomization
 - Higher variability reflecting the uncertainty of performance

Probability of stopping for futility

SCENARIO	P(Futility)
Negative	69.0%
No drug activity	27.9%
One arm mildly active	5.3%
One arm highly active	2.4%
Mild dose response	3.9%
High dose response	0.6%
Both arms mildly active	1.4%
Both arms highly active	0.3%

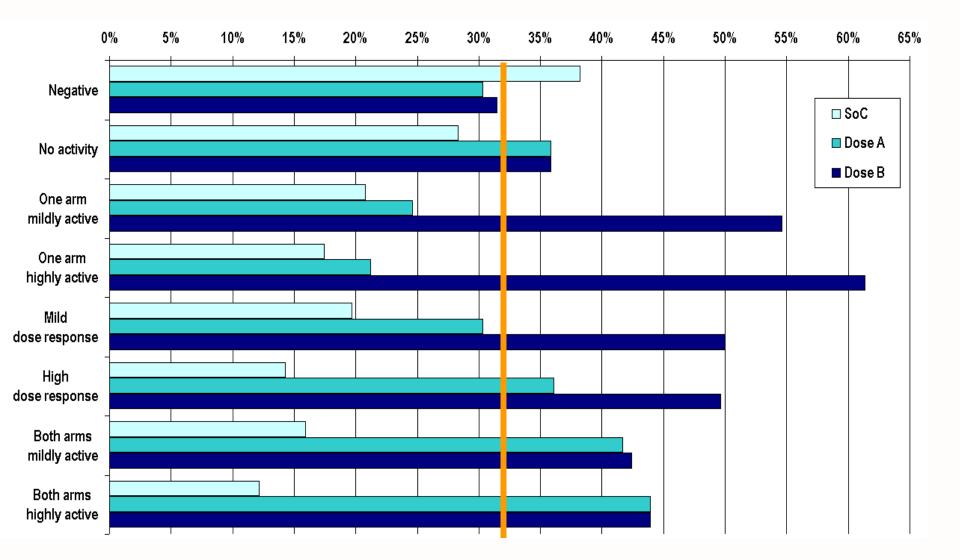
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Expected Number of Subjects Randomized Cute



% Number of Subjects Randomized without Burn-In Period

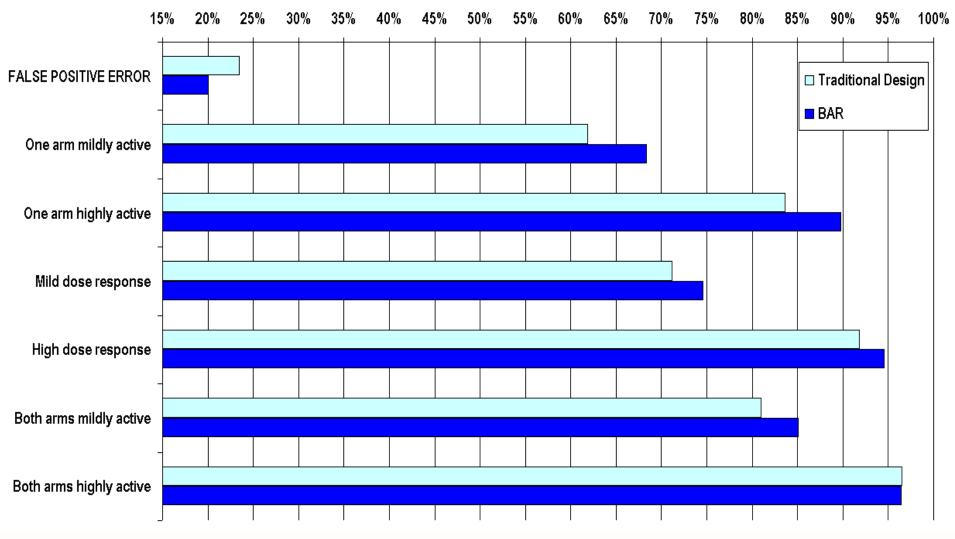






SCENARIO	π _{SoC}
Negative	57.1%
No drug activity	33.0%
One arm mildly active	10.3%
One arm highly active	4.8%
Mild dose response	8.8%
High dose response	2.6%
Both arms mildly active	4.7%
Both arms highly active	1.6%

Alpha and Power



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Discussion and Conclusions



- Bayesian Adaptive Randomization could be alternative design for Proof-Of-Concept studies in oncology
- Key points to consider when planning such designs:
 - Median Time to event / Recruitment rate
 - Schedule of assessment (PFS)
 - Control of covariates / Presence of treatment predictive factors
 - Model does not take into account safety
 - Operational burden (eCRF, blinding, etc)
- Simulations are of key importance to evaluate the applicability and the expected benefit of this design.



- **Cheung YK, Inoue LYT, Wathen JK, et al**. Continuous Bayesian adaptive randomization based on event times with covariates. StatMed 2006; 25:55-70
- **Berry D, Eick SG**. Adaptive assignment versus balanced randomization in clinical trials: a decision analysis. StatMed 1995; 14:231-246
- **Rosenberger WF**. New directions in adaptive designs. Statistical Science; 1996; 11:137-149
- Giles FJ, Kantarjian HM, Cortes JE, et al. Adaptive Randomized Study of Idarubicin and Cytarabine Versus Troxacitabine and Cytarabine Versus Troxacitabine and Idarubicin in Untreated Patients 50 Years or Older With Adverse Karyotype Acute Myeloid Leukemia. J Clin Oncol 21:1722-1727



THANK YOU!



BACKUP SLIDES



- Primary outcome is Progression Free Survival (PFS) time (measured in months)
- X_{ij} the theoretical PFS time for patient i on therapy j, where j = 1,2,3 and i = 1,2,...
- We assume that X_{ij} has an exponential distribution with median η_j and corresponding mean λ_i

$$f(x_{ij}) = \frac{1}{\lambda_j} e^{-x_{ij}/\lambda_j} \equiv \frac{\ln 2}{\eta_j} e^{-x_{ij}\ln 2/\eta_j}.$$

The model: Prior distributions



- The median parameters, η₁, η₂ and η₃ follow independent inverse gamma distributions with shape parameters α_j and scale parameters β_j
- The data for each patient consist of a pair of the form (Z_{ij}, δ_{ij}) where Z_{ij} is the observed PFS time for patient *i* under treatment *j*, and δ_{ij} is the indicator variable taking the value 1 if the event is observed and 0 if the patient is censored. The likelihood then becomes:

$$L(Z;\eta_j) = \left(\frac{\ln 2}{\eta_j}\right)^{E_j^+} \exp\left(-\frac{\ln(2) T_j^+}{\eta_j}\right)$$

• where $T_{j}^{+} = \sum_{i=1}^{n_{j}} Z_{ij}$ and $E_{j}^{+} = \sum_{i=1}^{n_{j}} \delta_{ij}$



- The combination of the exponential likelihood along with the Inverse Gamma priors result in the posterior distribution of the η_j parameters,
 - i.e., $\pi(\eta_j | \mathbb{Z})$ being independent Inverse Gamma distributions as well with
 - shape parameters $\alpha_j + E_j^+$ and scale parameters $\beta_j + \ln(2) T_j^+$

Choice of Priors



- The prior distribution of median PFS time (η_P) for the placebo arm chosen based on results of publications (see Small 2006) in the same setting, where the median PFS time for active treatment was 11.7 weeks (95% Cl, 9.1 to 16.6) and 10.0 weeks (95% Cl, 8.7 to 13.1) for placebo-treated patients.
- Hence, assumed a prior distribution of the median PFS time for placebo with an expected median PFS of 2.76 months (12weeks) and a 95% confidence interval of 1.86 to 4.09 months (8.06 to 17.7 weeks). This corresponds to:

 $\eta_P \sim IG(25.0836; 66.4708)$

 For the active treatments we assumed the same expected value but a higher variance in order to reflect the uncertainty over the drug.

 $\eta_{exp} \sim IG(17.2352; 44.809152)$