



# Reducing the Sample Size of Diagnostic-Biomarker-Validation Designs by a Bayesian Framework

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### **Outline**

#### Problem setting

Accuracy definition
Index definition
Incorporating pre-validation information

#### Bayesian framework

Development stage Validation stage Information transfer

#### Simulation study

Settings Results

#### Conclusions





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### Problem setting

# Develop a diagnostic biomarker-index and efficiently validate its accuracy

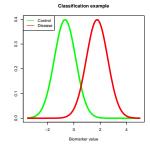
- Area under the Receiver Operating Characteristics (ROC) curve (AUC) as measure of accuracy
- Define index as linear combination of biomarkers maximizing AUC
- Incorporate information from development to validation stage

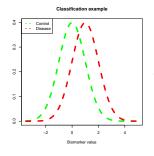
=> To this end a Bayesian framework will be proposed

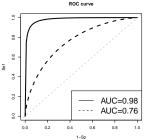




### Area under the Receiver Operating Characteristics curve











### Data assumptions and notation

#### Underlying true biomarker distribution

- Mixture of two K-variate normal distributions by true disease status (D)
  - $ightharpoonup |\mathbf{Y}|_{D=0} \sim N_{\mathcal{K}}(oldsymbol{\mu}_0, oldsymbol{\Sigma}_0)$
  - $ightharpoonup |\mathbf{Y}|_{D=1} \sim N_K(oldsymbol{\mu}_1, oldsymbol{\Sigma}_1)$
- Reference test (T) is imperfect
  - Se: Unknown sensitivity of the reference test
  - Sp: Unknown specificity of the reference test
  - Conditionally on true disease status, misclassification independent of biomarker value
- $\triangleright$   $\theta$ : Unknown true prevalence of disease in the data set
- Assume for the rest of the presentation K=3





### Definintion of biomarker-index

Linear combination maximizing AUC of the form\*:

$$egin{aligned} \mathbf{a'Y}|_{\mathit{D}=0} &\sim \mathit{N}(\mathbf{a'}\mu_0,\mathbf{a'}oldsymbol{\Sigma_0}\mathbf{a}) \ \mathbf{a'Y}|_{\mathit{D}=1} &\sim \mathit{N}(\mathbf{a'}\mu_1,\mathbf{a'}oldsymbol{\Sigma_1}\mathbf{a}) \end{aligned}$$

For which:

a' 
$$\propto (\mathbf{\Sigma}_0 + \mathbf{\Sigma}_1)^{-1} (\mu_1 - \mu_0)$$

Area Under the ROC Curve:

$$extit{AUC}_{ extit{Index}} = \Phi \left\{ \left[ (oldsymbol{\mu}_{ extsf{1}} - oldsymbol{\mu}_{ extsf{0}})'(oldsymbol{\Sigma}_{ extsf{0}} + oldsymbol{\Sigma}_{ extsf{1}})^{-1}(oldsymbol{\mu}_{ extsf{1}} - oldsymbol{\mu}_{ extsf{0}}) 
ight]^{rac{1}{2}} 
ight\}$$

\*Siu, J.Q., and Liu, J.S. (1993)

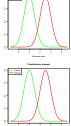




### **Development**

#### Data





#### Results









Incorporating pre-validation information

### Development

### **Validation**

#### Data





### <u>Results</u>





### <u>Data</u>







### <u>Results</u>



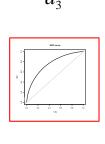




Incorporating pre-validation information

### **Validation**





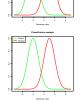
Results

# Data





Results







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### Bayesian latent-class mixture model

#### Full data likelihood

$$\begin{split} &L(\boldsymbol{\mu}_0, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_0, \boldsymbol{\Sigma}_1, \boldsymbol{\theta}, Se, Sp|\mathbf{Y}, \mathbf{T}, \mathbf{D}) \\ &= \prod_{i=1}^N \left( \boldsymbol{\theta} Se^{t_i} (1 - Se)^{(1-t_i)} \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}_1|}} \times \textit{EXP} \left\{ -\frac{1}{2} \left( \mathbf{Y}_i - \boldsymbol{\mu}_1 \right)' \boldsymbol{\Sigma}_1^{-1} \left( \mathbf{Y}_i - \boldsymbol{\mu}_1 \right) \right\} \right)^{d_i} \\ &\times \left( (1 - \boldsymbol{\theta}) (1 - Sp)^{t_i} Sp^{(1-t_i)} \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}_0|}} \times \textit{EXP} \left\{ -\frac{1}{2} \left( \mathbf{Y}_i - \boldsymbol{\mu}_0 \right)' \boldsymbol{\Sigma}_0^{-1} \left( \mathbf{Y}_i - \boldsymbol{\mu}_0 \right) \right\} \right)^{(1-d_i)} \end{split}$$



└ Development stage

### **Prior distributions**

#### Hyperprior

 $\theta \sim \text{Uniform}(0.1,0.9)$ 

#### **Priors**

$$D_i \sim \text{Bernoulli}(\theta)$$

Se = Sp  $\sim$  Beta(1,1)T(0.51, $\infty$ )

(Observation i: 1,...,N)





### **Prior distributions**

Set 
$$\Sigma_i = V_i R_i V_i^*$$

For:  $V_j = \sigma_{k,j}I_3$  and  $R_j$  is a correlation matrix.

Then: 
$$\mathbf{C}_j = \begin{pmatrix} 1 & c_{j,12} & c_{j,13} \\ 0 & c_{j,22} & c_{j,23} \\ 0 & 0 & c_{j,33} \end{pmatrix}$$
 = Cholesky factor of  $\mathbf{R}_j$ .

$$\sigma_{k,j} \sim \mathsf{Uniform}(\mathsf{0,1000})$$

$$c_{j,12} = \rho_{j,12} \sim \text{Uniform(-1,1)}$$
 $c_{j,13} = \rho_{j,13} \sim \text{Uniform(-1,1)}$ 
 $c_{j,23} \sim \text{Uniform}\left(-\sqrt{1-\rho_{j,13}^2}, \sqrt{1-\rho_{j,13}^2}\right)$ 

$$\rho_{j,23} = \rho_{j,12}\rho_{j,13} + c_{j,22}c_{j,23}$$

\* Wei, Y and Higgins, J.P.T (2013)





### **Prior distributions**

$$\textit{AUC}_{\textit{Index}} = \Phi \left\{ \left[ (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_0)' (\boldsymbol{\Sigma}_0 + \boldsymbol{\Sigma}_1)^{-1} (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_0) \right]^{\frac{1}{2}} \right\}$$

#### Reparameterize:

$$egin{aligned} &AUC_{\textit{Index}} = \Phi \left\{ \sqrt{m{\Delta}' m{\Delta}} 
ight\} \ & ext{Where } m{\Delta} = m{L}(\mu_1 - \mu_0) \ & ext{For } m{L} = ext{the Cholesky factor of } (m{\Sigma}_0 + m{\Sigma}_1)^{-1} \end{aligned}$$

#### **Priors**

$$egin{aligned} oldsymbol{\Delta} &\sim \emph{N}_{3}(\kappa, \Psi) \ \mu_{0\emph{k}} &\sim \emph{N}(0, 10^{6}) \ \ (\emph{k: 1,...,3}) \ oldsymbol{\mu}_{1} &= oldsymbol{\Delta} \emph{L}^{-1} + oldsymbol{\mu}_{0} \end{aligned}$$





### Bayesian latent-class mixture model

$$\mathbf{Y}_{\mathit{Index}} = \hat{\mathbf{a}}' \mathbf{Y}_{\mathit{Val}}$$

(Biomarker index observations)

#### Full data likelihood

$$\begin{split} &L(\mu_0, \mu_1, \sigma_0, \sigma_1, \theta, \textit{Se}, \textit{Sp}|\mathbf{Y}_{\textit{Index}}, \mathbf{T}_{\textit{Val}}, \mathbf{D}_{\textit{Val}}) \\ &= \prod_{i=1}^{\textit{N}} \left(\theta \textit{Se}^{\textit{I}_{\textit{Val}_i}} (1 - \textit{Se})^{(1 - \textit{I}_{\textit{Val}_i})} \frac{1}{\sqrt{2\pi\sigma_1^2}} \times \textit{EXP}\left\{-\frac{(Y_{\textit{Index}_i} - \mu_1)^2}{\sigma_1^2}\right\}\right)^{\textit{d}_{\textit{Val}_i}} \\ &\times \left((1 - \theta)(1 - \textit{Sp})^{\textit{I}_{\textit{Val}_i}} \textit{Sp}^{(1 - \textit{I}_{\textit{Val}_i})} \frac{1}{\sqrt{2\pi\sigma_0^2}} \times \textit{EXP}\left\{-\frac{(Y_{\textit{Index}_i} - \mu_0)^2}{\sigma_0^2}\right\}\right)^{(1 - \textit{d}_{\textit{Val}_i})} \end{split}$$





└─ Validation stage

### Prior distributions

#### **Hyperprior**

 $\theta \sim \text{Uniform}(0.1,0.9)$ 

#### **Priors**

 $D_i \sim \text{Bernoulli}(\theta)$ 

(Observation i: 1,...,N)

Se = Sp  $\sim$  Beta(1,1)T(0.51, $\infty$ )  $\sigma_i \sim$  Uniform(0,1000)

[j:0,1]





### **Prior distributions**

$$extit{AUC}_{ extit{Index}} = \Phi \left\{ rac{(\mu_1 - \mu_0)}{\sqrt{\sigma_0^2 + \sigma_1^2}} 
ight\}$$

#### Reparameterize:

$$AUC_{Index} = \Phi \left\{ \gamma \right\}$$
 Where  $\gamma = \frac{(\mu_1 - \mu_0)}{\sqrt{\sigma_0^2 + \sigma_1^2}}$ 

#### **Priors**

$$\gamma \sim N(\lambda, \tau^2)$$

$$\mu_0 \sim N(0, 10^6)$$

$$\mu_1 = \gamma \times \sqrt{\sigma_0^2 + \sigma_1^2} + \mu_0$$



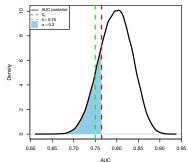


### Validation criterion

#### Based on Bayesian hypothesis testing paradigm:

 $H_0: AUC \leq \delta$  $H_1: AUC > \delta$ 





Consider result significant when posterior probability of AUC exceeding  $\delta$  is larger than 1  $-\alpha$ .





### Incorporate pre-validation information

$$\begin{array}{ll} \textbf{Development} & \textbf{Validation} \\ \phi^{-1} \left\{ \textit{AUC}_{\textit{Index}} \right\} = \sqrt{\Delta' \Delta} & \approx & \gamma = \phi^{-1} \left\{ \textit{AUC}_{\textit{Index}} \right\} \end{array}$$

Take approximation to posterior distribution of  $\sqrt{\Delta'\Delta}$  as prior distribution for  $\gamma$ 





### Incorporate pre-validation information

$$\begin{array}{ll} \textbf{Development} & \textbf{Validation} \\ \phi^{-1} \left\{ \textit{AUC}_{\textit{Index}} \right\} = \sqrt{\Delta' \Delta} & \approx & \gamma = \phi^{-1} \left\{ \textit{AUC}_{\textit{Index}} \right\} \end{array}$$

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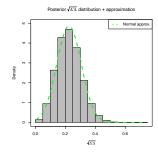


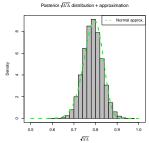
## $\sqrt{\Delta'\Delta}$ posterior approximation

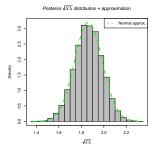
Prior:  $\gamma \sim N(\lambda, \tau^2)$ 

Where 
$$\lambda=\overline{\mathbf{x}}_{\sqrt{\Delta'\Delta}_{1:M}}$$
, and  $\tau^2=\mathbf{s}_{\sqrt{\Delta'\Delta}_{1:M}}^2$ 

(For M MCMC samples)











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#### **GOAL**

# Establish difference in power to validate AUC of biomarker index when ignoring vs incorporating pre-validation information

#### For 3 correlated biomarkers

$$\theta = 0.5$$

$$Se = Sp = 0.85$$

Mixture component parameters set such that:

AUC of biomarker 1 = 0.75

AUC of biomarker 2 = 0.75

AUC of biomarker 3 = 0.75

$$AUC_{Index} = 0.78$$





### **Development Stage**

- $N_{Dev} = 400$
- $\hat{\mathbf{a}}'$  = posterior median of  $\mathbf{a}'$

#### Validation Stage

- ightharpoonup 200 data sets for  $N_{Val} = 100, 400, 600, and 800$
- Power = proportion of simulations for which P(AUC > 0.75|data) > 0.80
- Prior AUC information:

- ► Ignoring AUC information (Red)
  - Prior  $\gamma: N(0,1)$
- ► Incorporating AUC information (Blue)
  - Prior  $\gamma$ : Discounted normal approx. to posterior predictive distribution of  $\sqrt{\Delta'\Delta}$





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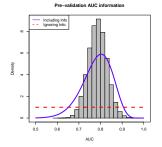


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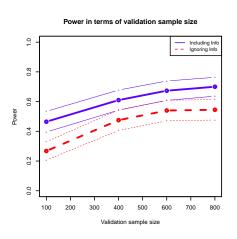


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### Simulated Power

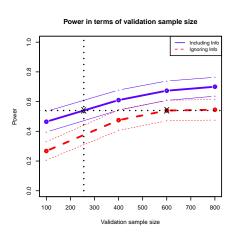


- Increasing validation sample size increases power
- Incorporating pre-validation information significantly increases power
- ▶ Reduction of about ½ of sample size to maintain power





### Simulated Power



- Increasing validation sample size increases power
- Incorporating pre-validation information significantly increases power
- Reduction of about <sup>1</sup>/<sub>2</sub> of sample size to maintain power





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#### Conclusions

- Bayesian framework:
  - Allows including pre-validation information into validation stage
    - By approximating posterior AUC information as prior
    - Also other pre-validation information possible
- Simulation study
  - Power to reach validation is significantly increased
  - Sample size reduction for equal power





### Further considerations

- Other validation criteria
- Robustness to miss-estimated linear combination coefficients
- Extend to incorporate non-normally distributed biomarkers
- Evaluate impact of conditional independence assumption





#### References

- Su, J.Q., Liu, J.S.: Linear combinations of multiple diagnostic markers.
   Journal of the American Statistical Association. 88, 1350–1355 (1993)
- Wei, Y, Higgins, P.T.: Bayesian multivariate meta-analysis with multiple outcomes. Statistics in Medicine (2013) doi: 10.1002/sim.5745

- Conclusions





Thank you for your attention!