



Universität  
Zürich<sup>UZH</sup>

## The quantile probability model

# A novel approach to Bayesian variable selection for clinical prediction

Rachel Heyard and Leonhard Held

Department of Biostatistics - EBPI - University of Zurich

22.06.2018

# Clinical prediction models

Clinical prediction models used for decision making

# Clinical prediction models

Clinical prediction models used for decision making

⇒ to identify low and high risk patients for a specific disease /  
pathogen / event /...

# Clinical prediction models

Clinical prediction models used for decision making

- ⇒ to identify low and high risk patients for a specific disease / pathogen / event /...
- ⇒ to better target treatment and prevention

# Clinical prediction models

Clinical prediction models used for decision making

- ⇒ to identify low and high risk patients for a specific disease / pathogen / event /...
- ⇒ to better target treatment and prevention
- ⇒ to improve medical practice

# Objective Bayesian variable selection for GLMs

Select among  $p$  variables  $\rightarrow 2^p$  candidate models  $\mathcal{M}_j$

---

<sup>1</sup> Hu and Johnson. *Bayesian model selection using test statistics*. Journal of the Royal Statistical Society, 2009

<sup>2</sup> Held, Sabanés Bové and Gravestock. *Approximate Bayesian model selection with the deviance statistic*. Statistical Science, 2015

# Objective Bayesian variable selection for GLMs

Select among  $p$  variables  $\rightarrow 2^p$  candidate models  $\mathcal{M}_j$

**Posterior model probability:**

$$\Pr(\mathcal{M}_j \mid \text{data}) = \frac{\text{TBF}_{j,0} \Pr(\mathcal{M}_j)}{\sum_{k \in \mathcal{J}} \text{TBF}_{k,0} \Pr(\mathcal{M}_k)}$$

---

<sup>1</sup> Hu and Johnson. *Bayesian model selection using test statistics*. Journal of the Royal Statistical Society, 2009

<sup>2</sup> Held, Sabanés Bové and Gravestock. *Approximate Bayesian model selection with the deviance statistic*. Statistical Science, 2015

# Objective Bayesian variable selection for GLMs

Select among  $p$  variables  $\rightarrow 2^p$  candidate models  $\mathcal{M}_j$

**Posterior model probability:**

$$\Pr(\mathcal{M}_j \mid \text{data}) = \frac{\text{TBF}_{j,0} \Pr(\mathcal{M}_j)}{\sum_{k \in \mathcal{J}} \text{TBF}_{k,0} \Pr(\mathcal{M}_k)}$$

Generalized  $g$ -prior<sup>1,2</sup> leads to

$$\text{TBF}_{j,0} = \frac{p(z_j \mid \mathcal{M}_j)}{p(z_j \mid \mathcal{M}_0)} = (g + 1)^{-d_j/2} \exp\left(\frac{g}{g + 1} \frac{z_j}{2}\right)$$

$z_j$  is the deviance statistic of model  $\mathcal{M}_j$  with  $d_j$  degrees of freedom

---

<sup>1</sup> Hu and Johnson. *Bayesian model selection using test statistics*. Journal of the Royal Statistical Society, 2009

<sup>2</sup> Held, Sabanés Bové and Gravestock. *Approximate Bayesian model selection with the deviance statistic*. Statistical Science, 2015



# Objective Bayesian variable selection for GLMs

**Posterior inclusion probability** of  $x_k$ , with  $k \in \{1, \dots, p\}$ :

$$\Pr(x_k \in \mathcal{M} \mid \text{data}) = \sum_{j \in \mathcal{J}: x_k \in \mathcal{M}_j} \Pr(\mathcal{M}_j \mid \text{data})$$

# Objective Bayesian variable selection for GLMs

**Posterior inclusion probability** of  $x_k$ , with  $k \in \{1, \dots, p\}$ :

$$\Pr(x_k \in \mathcal{M} \mid \text{data}) = \sum_{j \in \mathcal{J}: x_k \in \mathcal{M}_j} \Pr(\mathcal{M}_j \mid \text{data})$$

**Median probability model (MPM)** <sup>3</sup>:

Include  $x_k$  if  $\Pr(x_k \in \mathcal{M} \mid \text{data}) \geq 0.5$

---

<sup>3</sup>Barbieri and Berger. *Optimal predictive model selection*. Annals of Statistics, 2004

# Objective Bayesian variable selection for GLMs

**Posterior inclusion probability** of  $x_k$ , with  $k \in \{1, \dots, p\}$ :

$$\Pr(x_k \in \mathcal{M} \mid \text{data}) = \sum_{j \in \mathcal{J}: x_k \in \mathcal{M}_j} \Pr(\mathcal{M}_j \mid \text{data})$$

**Median probability model (MPM)** <sup>3</sup>:

Include  $x_k$  if  $\Pr(x_k \in \mathcal{M} \mid \text{data}) \geq 0.5$

$\Rightarrow$  MPM = optimal for prediction under the squared error loss when selecting among normal linear models

---

<sup>3</sup>Barbieri and Berger. *Optimal predictive model selection*. Annals of Statistics, 2004

# Objective Bayesian variable selection for GLMs

**Posterior inclusion probability** of  $x_k$ , with  $k \in \{1, \dots, p\}$ :

$$\Pr(x_k \in \mathcal{M} \mid \text{data}) = \sum_{j \in \mathcal{J}: x_k \in \mathcal{M}_j} \Pr(\mathcal{M}_j \mid \text{data})$$

**Median probability model (MPM)** <sup>3</sup>:

Include  $x_k$  if  $\Pr(x_k \in \mathcal{M} \mid \text{data}) \geq 0.5$

$\Rightarrow$  MPM = optimal for prediction under the squared error loss when selecting among normal linear models

**BUT** MPM also highly sensitive to prior choices

---

<sup>3</sup>Barbieri and Berger. *Optimal predictive model selection*. Annals of Statistics, 2004

# Prior choices

---

<sup>4</sup>Scott and Berger. *Bayes and empirical-Bayes multiplicity adjustment in the variable-selection problem*. Annals of Statistics, 2010

# Prior choices

**Model prior** of  $\mathcal{M}_j$  with  $p_j$  variables

---

<sup>4</sup>Scott and Berger. *Bayes and empirical-Bayes multiplicity adjustment in the variable-selection problem*. Annals of Statistics, 2010

# Prior choices

**Model prior** of  $\mathcal{M}_j$  with  $p_j$  variables

⇒ fixed **prior inclusion probability** at  $q = 1/2$

→ uniform prior on the model space:  $\Pr(\mathcal{M}_j) = 1/2^p$

---

<sup>4</sup>Scott and Berger. *Bayes and empirical-Bayes multiplicity adjustment in the variable-selection problem*. Annals of Statistics, 2010

# Prior choices

**Model prior** of  $\mathcal{M}_j$  with  $p_j$  variables

⇒ fixed **prior inclusion probability** at  $q = 1/2$

→ uniform prior on the model space:  $\Pr(\mathcal{M}_j) = 1/2^p$

⇒ beta distribution on the **prior inclusion probability**  $q \sim \text{Be}(a, b)$

$$\rightarrow \Pr(\mathcal{M}_j \mid a, b) = \frac{B(a + p_j, b + p - p_j)}{B(a, b)}$$

---

<sup>4</sup>Scott and Berger. *Bayes and empirical-Bayes multiplicity adjustment in the variable-selection problem*. Annals of Statistics, 2010



# Prior choices

**Model prior** of  $\mathcal{M}_j$  with  $p_j$  variables

⇒ fixed **prior inclusion probability** at  $q = 1/2$

→ uniform prior on the model space:  $\Pr(\mathcal{M}_j) = 1/2^p$

⇒ beta distribution on the **prior inclusion probability**  $q \sim \text{Be}(a, b)$

$$\rightarrow \Pr(\mathcal{M}_j \mid a, b) = \frac{B(a + p_j, b + p - p_j)}{B(a, b)}$$

$a = b = 1 \rightarrow$  multiplicity-corrected model prior <sup>4</sup>

---

<sup>4</sup>Scott and Berger. *Bayes and empirical-Bayes multiplicity adjustment in the variable-selection problem*. Annals of Statistics, 2010

# Prior choices

**Model prior** of  $\mathcal{M}_j$  with  $p_j$  variables

⇒ fixed **prior inclusion probability** at  $q = 1/2$

→ uniform prior on the model space:  $\Pr(\mathcal{M}_j) = 1/2^p$

⇒ beta distribution on the **prior inclusion probability**  $q \sim \text{Be}(a, b)$

$$\rightarrow \Pr(\mathcal{M}_j \mid a, b) = \frac{B(a + p_j, b + p - p_j)}{B(a, b)}$$

$a = b = 1 \rightarrow$  multiplicity-corrected model prior <sup>4</sup>

## Definition of **g**

---

<sup>4</sup>Scott and Berger. *Bayes and empirical-Bayes multiplicity adjustment in the variable-selection problem*. Annals of Statistics, 2010

# Prior choices

**Model prior** of  $\mathcal{M}_j$  with  $p_j$  variables

⇒ fixed **prior inclusion probability** at  $q = 1/2$

→ uniform prior on the model space:  $\Pr(\mathcal{M}_j) = 1/2^p$

⇒ beta distribution on the **prior inclusion probability**  $q \sim \text{Be}(a, b)$

$$\rightarrow \Pr(\mathcal{M}_j \mid a, b) = \frac{B(a + p_j, b + p - p_j)}{B(a, b)}$$

$a = b = 1 \rightarrow$  multiplicity-corrected model prior <sup>4</sup>

## Definition of $g$

⇒ Local empirical Bayes:  $\hat{g} = \max\{z_j/d_j - 1, 0\}$

⇒ hyper- $g/n$ :  $\frac{g/n}{g/n+1} \sim U(0, 1)$

⇒ Zellner-Siow adapted

---

<sup>4</sup>Scott and Berger. *Bayes and empirical-Bayes multiplicity adjustment in the variable-selection problem*. Annals of Statistics, 2010

# Case study

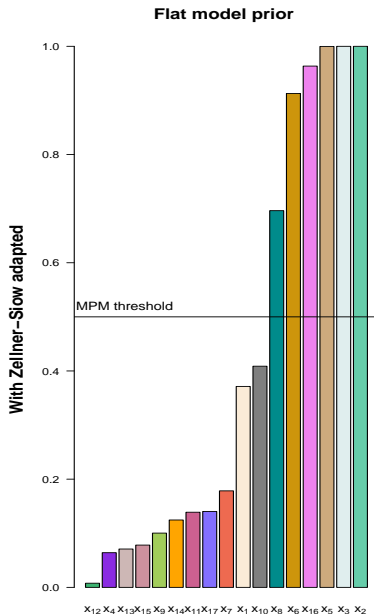
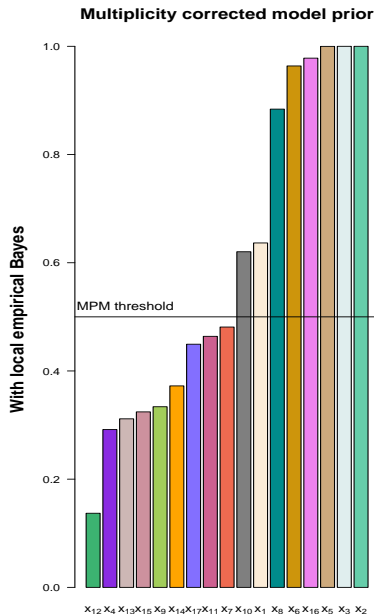
Application on **GUSTO-I** data using logistic regression

- ⇒ Randomized study for comparison of 4 different treatments in over 40'000 acute myocardial infarction patients
- ⇒ We use a publicly available subgroup from the Western region of the USA ( $n = 2188$ )
- ⇒ prediction of **30-days survival**

We focus on **17 covariates**

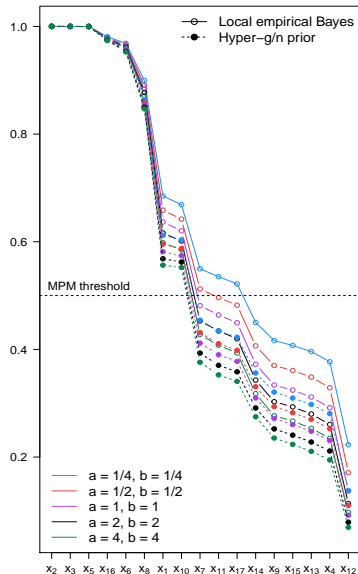
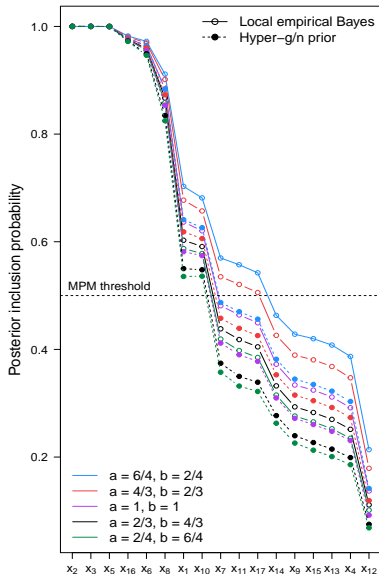
# Sensitivity of the MPM

with respect to the definition of  $g$  and the model prior



# Sensitivity of the MPM

with respect to the hyper-parameters  $a$  and  $b$



# The quantile probability model

# The quantile probability model

- Define  $p + 1$  candidate models  $\{\mathcal{M}_0, \mathcal{M}_1, \dots, \mathcal{M}_p\}$  based on inclusion probabilities



# The quantile probability model

- Define  $p + 1$  candidate models  $\{\mathcal{M}_0, \mathcal{M}_1, \dots, \mathcal{M}_p\}$  based on inclusion probabilities
- Compute predictive model selection criterion:  
DIC, WAIC or LOO-IC

# The quantile probability model

- Define  $p + 1$  candidate models  $\{\mathcal{M}_0, \mathcal{M}_1, \dots, \mathcal{M}_p\}$  based on inclusion probabilities
- Compute predictive model selection criterion:  
DIC, WAIC or LOO-IC
- Select the model with lowest information criterion  $\rightarrow$  QPM

# Predictive model selection information criteria (IC)

The **deviance information criterion** (DIC):

$$-2 \log f(\mathbf{y} \mid \hat{\theta}) + 2p_{\text{DIC}}$$

---

<sup>5</sup>Gelman, Hwang and Vehtari. *Understanding predictive information criteria for Bayesian models*. Statistics and Computing, 2014

# Predictive model selection information criteria (IC)

The **deviance information criterion** (DIC):

$$-2 \log f(\mathbf{y} \mid \hat{\theta}) \quad + \quad 2p_{\text{DIC}}$$

[adjustment of the log pred. density]    +    [approximate bias correction]

---

<sup>5</sup>Gelman, Hwang and Vehtari. *Understanding predictive information criteria for Bayesian models*. Statistics and Computing, 2014

# Predictive model selection information criteria (IC)

The **deviance information criterion** (DIC):

$$-2 \log f(\mathbf{y} \mid \hat{\theta}) + 2p_{\text{DIC}}$$

[adjustment of the log pred. density] + [approximate bias correction]

Similarly, **Watanabe-Akaike** and **leave-one-out cross-validation IC** (WAIC and LOO-IC) can be defined <sup>5</sup>

---

<sup>5</sup>Gelman, Hwang and Vehtari. *Understanding predictive information criteria for Bayesian models*. Statistics and Computing, 2014

# Predictive model selection information criteria (IC)

The **deviance information criterion** (DIC):

$$-2 \log f(\mathbf{y} \mid \hat{\theta}) + 2p_{\text{DIC}}$$

[adjustment of the log pred. density] + [approximate bias correction]

Similarly, **Watanabe-Akaike** and **leave-one-out cross-validation IC** (WAIC and LOO-IC) can be defined <sup>5</sup>

→ computation possible with `loo`-package in R

---

<sup>5</sup>Gelman, Hwang and Vehtari. *Understanding predictive information criteria for Bayesian models*. Statistics and Computing, 2014

# Predictive model selection information criteria (IC)

The **deviance information criterion** (DIC):

$$\begin{array}{rcl} -2 \log f(\mathbf{y} \mid \hat{\theta}) & + & 2p_{\text{DIC}} \\ \text{[adjustement of the log pred. density]} & + & \text{[approximate bias correction]} \end{array}$$

Similarly, **Watanabe-Akaike** and **leave-one-out cross-validation IC** (WAIC and LOO-IC) can be defined <sup>5</sup>

→ computation possible with `loo`-package in R

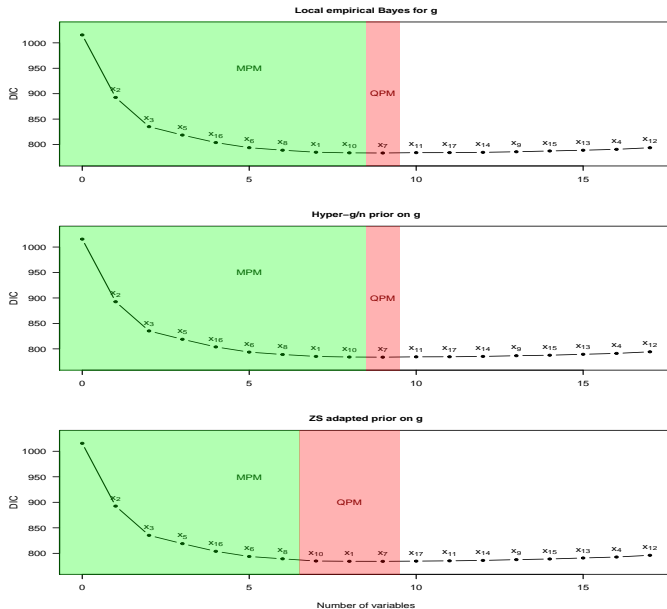
→ IC does not depend on model prior !

---

<sup>5</sup>Gelman, Hwang and Vehtari. *Understanding predictive information criteria for Bayesian models*. Statistics and Computing, 2014

# The quantile probability model

QPM model  $M_9$  independent from the definition of  $g$





# Monte Carlo error of DIC

Sampling from the posterior distribution → **Monte Carlo** (MC) **error**

# Monte Carlo error of DIC

Sampling from the posterior distribution → **Monte Carlo** (MC) **error**

## **Brute force approach:**

Replicate the DIC calculation  $N$  times and compute sample standard deviation of the  $N$  DIC estimates

# Monte Carlo error of DIC

Sampling from the posterior distribution → **Monte Carlo** (MC) **error**

## **Brute force approach:**

Replicate the DIC calculation  $N$  times and compute sample standard deviation of the  $N$  DIC estimates

⇒ very time-consuming

# New efficient approach for Monte Carlo error of DIC

# New efficient approach for Monte Carlo error of DIC

Suppose DIC has been computed using MC sample of size  $S$

# New efficient approach for Monte Carlo error of DIC

Suppose DIC has been computed using MC sample of size  $S$

A solid orange rectangular box containing the text 'DIC'.

DIC

# New efficient approach for Monte Carlo error of DIC

Suppose DIC has been computed using MC sample of size  $S$



# New efficient approach for Monte Carlo error of DIC

Suppose DIC has been computed using MC sample of size  $S$





# New efficient approach for Monte Carlo error of DIC

Suppose DIC has been computed using MC sample of size  $S$



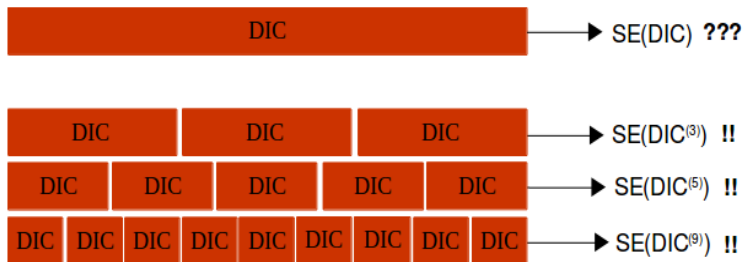
# New efficient approach for Monte Carlo error of DIC

Suppose DIC has been computed using MC sample of size  $S$



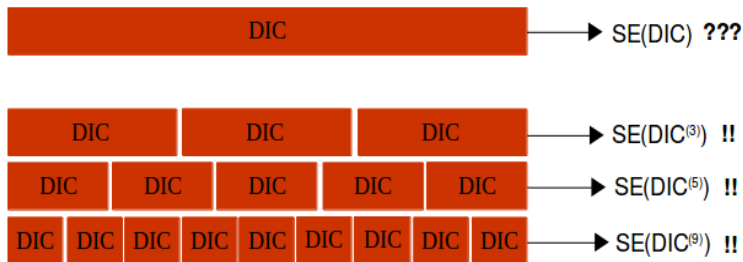
# New efficient approach for Monte Carlo error of DIC

Suppose DIC has been computed using MC sample of size  $S$



# New efficient approach for Monte Carlo error of DIC

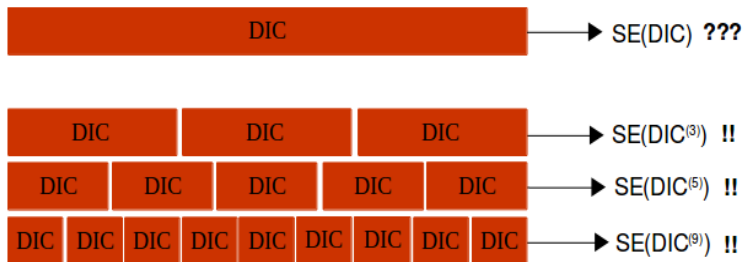
Suppose DIC has been computed using MC sample of size  $S$



with the **square-root** law we obtain  $SE(DIC^{(k)}) \propto \sqrt{\frac{1}{S/k}} \propto \sqrt{k}$

# New efficient approach for Monte Carlo error of DIC

Suppose DIC has been computed using MC sample of size  $S$



with the **square-root** law we obtain  $SE(DIC^{(k)}) \propto \sqrt{\frac{1}{S/k}} \propto \sqrt{k}$

$$\Rightarrow SE(DIC^{(k)}) = c \cdot \sqrt{k} \quad \Rightarrow \quad SE(DIC^{(1)}) = c$$

estimate  $c$  with weighted linear regression of  $SE(DIC^{(k)})$  on  $\sqrt{k}$  with weights  $k$

# Monte Carlo error of DIC

Application to case study

Case study:  $S = 50,000$  and  $k \in \{80, 100, 125, 160, 200\}$

# Monte Carlo error of DIC

Application to case study

Case study:  $S = 50,000$  and  $k \in \{80, 100, 125, 160, 200\}$

The **Monte Carlo standard error** of  $\Delta\text{DIC}_j = \text{DIC}_{j-1} - \text{DIC}_j$ :

$$\text{SE}(\Delta\text{DIC}_j) = \sqrt{\text{SE}(\text{DIC}_{j-1})^2 - \text{SE}(\text{DIC}_j)^2}$$

# Monte Carlo error of DIC

Application to case study

Case study:  $S = 50,000$  and  $k \in \{80, 100, 125, 160, 200\}$

The **Monte Carlo standard error** of  $\Delta\text{DIC}_j = \text{DIC}_{j-1} - \text{DIC}_j$ :

$$\text{SE}(\Delta\text{DIC}_j) = \sqrt{\text{SE}(\text{DIC}_{j-1})^2 - \text{SE}(\text{DIC}_j)^2}$$

	LEB		hyper- $g/n$		ZS adapted	
	$\Delta\text{DIC}_j$	(SE)	$\Delta\text{DIC}_j$	(SE)	$\Delta\text{DIC}_j$	(SE)
			$\vdots$			
$\mathcal{M}_8$ vs. $\mathcal{M}_9$	0.217	(0.031)	0.166	(0.034)	0.057	(0.033)
$\mathcal{M}_9$ vs. $\mathcal{M}_{10}$	-0.295	(0.031)	-0.342	(0.034)	-0.170	(0.032)
			$\vdots$			



# Monte Carlo error of DIC

Application to case study

Case study:  $S = 50,000$  and  $k \in \{80, 100, 125, 160, 200\}$

The **Monte Carlo standard error** of  $\Delta\text{DIC}_j = \text{DIC}_{j-1} - \text{DIC}_j$ :

$$\text{SE}(\Delta\text{DIC}_j) = \sqrt{\text{SE}(\text{DIC}_{j-1})^2 - \text{SE}(\text{DIC}_j)^2}$$

	LEB		hyper- $g/n$		ZS adapted	
	$\Delta\text{DIC}_j$	(SE)	$\Delta\text{DIC}_j$	(SE)	$\Delta\text{DIC}_j$	(SE)
			$\vdots$			
$\mathcal{M}_8$ vs. $\mathcal{M}_9$	0.217	(0.031)	0.166	(0.034)	0.057	(0.033)
$\mathcal{M}_9$ vs. $\mathcal{M}_{10}$	-0.295	(0.031)	-0.342	(0.034)	-0.170	(0.032)
			$\vdots$			

Conclusion: The DIC of the QPM  $\mathcal{M}_9$  is the unique minimum !

# Conclusion

- Drastically reduced candidate model space:  $2^p$  to  $p + 1$

# Conclusion

- Drastically reduced candidate model space:  $2^p$  to  $p + 1$
- In the application: QPM not sensitive with respect to prior choices

# Conclusion

- Drastically reduced candidate model space:  $2^p$  to  $p + 1$
- In the application: QPM not sensitive with respect to prior choices
- Model prior does not affect the information criterion

# Conclusion

- Drastically reduced candidate model space:  $2^p$  to  $p + 1$
- In the application: QPM not sensitive with respect to prior choices
- Model prior does not affect the information criterion
- Efficient method introduced to compute Monte Carlo error of IC

# Conclusion

- Drastically reduced candidate model space:  $2^p$  to  $p + 1$
  - In the application: QPM not sensitive with respect to prior choices
  - Model prior does not affect the information criterion
  - Efficient method introduced to compute Monte Carlo error of IC
- 
- Work revised for *Journal of Computational Statistics and Data Analysis*

***Thank you***

Questions? Comments?